

Multivariate Statistical Analysis

多元统计分析

2026年3月24日

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

3.1 协方差

- ▶ 协方差是相关性的一种度量.
- ▶ 协方差仅度量线性相关程度.
- ▶ 协方差与度量单位有关.
- ▶ 协方差为零也可能存在非线性相关.
- ▶ 协方差为零不一定独立.
- ▶ 独立协方差一定为零.
- ▶ 协方差为负对应着向下倾斜的散点图. 协方差为正对应着向上倾斜的散点图.
- ▶ 变量与自身的协方差就是方差 $\text{Cov}(X, X) = \sigma_{XX} = \sigma_X^2$.
- ▶ 对于较小的 n , 在计算协方差时, 我们用 $\frac{1}{n-1}$ 代替 $\frac{1}{n}$.

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

3.2 相关系数

- ▶ 相关系数是标准化之后对相关性的—种度量.
- ▶ 相关系数的绝对值不超过 1.
- ▶ 相关系数仅度量线性相关程度.
- ▶ 不相关亦可能存在非线性的相关.
- ▶ 不相关不一定独立. 但对两个正态分布的变量而言, 不相关则意味着独立.
- ▶ 独立—定不相关.
- ▶ 负相关对应着向下倾斜的散点图. 正相关对应着向上倾斜的散点图.
- ▶ Fisher 的 Z 变换 $W = \frac{1}{2} \log \left(\frac{1 + r_{XY}}{1 - r_{XY}} \right)$ 可用于相关系数的假设检验.
- ▶ 对小样本, 可用修正后的 Fisher 的 Z 变换 $W^* = W - \frac{3W + \tanh(W)}{4(n - 1)}$.

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

3.3 描述性统计

- ▶ 数据矩阵的重心由其均值向量 $\bar{\mathbf{x}} = \frac{1}{n} \mathcal{X}^T \mathbf{1}_n$ 给出.

- ▶ 数据矩阵观测值的离散程度由样本协方差矩阵给出:

$$\begin{aligned}
 \mathcal{S} &= \frac{1}{n} \sum_{k=1}^n (\mathbf{x}_k - \bar{\mathbf{x}}) (\mathbf{x}_k - \bar{\mathbf{x}})^T \\
 \mathcal{S}_u = \frac{n}{n-1} \mathcal{S} & \text{ 是总体协方差矩阵 } \Sigma \text{ 的无偏估计} \\
 &= \frac{1}{n} \mathcal{X}^T \mathcal{X} - \bar{\mathbf{x}} \bar{\mathbf{x}}^T \\
 &= \frac{1}{n} \mathcal{X}^T \mathcal{H} \mathcal{X}
 \end{aligned}$$

样本协方差矩阵 \mathcal{S} 非负定: $\mathcal{S} \geq 0$

中心化矩阵: $\mathcal{H} = \mathcal{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^T$
 $\mathcal{H}^T = \mathcal{H}, \quad \mathcal{H}^2 = \mathcal{H}$

- ▶ 样本相关矩阵由 $\mathcal{R} = \mathcal{D}^{-1/2} \mathcal{S} \mathcal{D}^{-1/2}$ 给出.

对角阵: $\mathcal{D}^{-1/2} = \text{diag} \left(s_{X_1 X_1}^{-1/2}, s_{X_2 X_2}^{-1/2}, \dots, s_{X_p X_p}^{-1/2} \right)$

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

- 线性变换: $Y = \mathcal{A}_{q \times p} X$

变换后的数据矩阵: $y = \mathcal{X} \mathcal{A}^T$

初始数据矩阵

$$\bar{y} = \mathcal{A} \bar{x}$$

$$\mathcal{S}_y = \mathcal{A} \mathcal{S}_x \mathcal{A}^T$$

- Mahalanobis 变换: $z_i = \mathcal{S}^{-1/2} (x_i - \bar{x})$, $i = 1, 2, \dots, n$

变换后的数据矩阵: $\mathcal{Z} = (z_1, z_2, \dots, z_n)^T = \mathcal{H} \mathcal{X} \mathcal{S}^{-1/2}$

$\mathcal{S}_{\mathcal{Z}} = \mathcal{I}_p$, 消除了变量之间的相关性

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

- 简单线性回归模型

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, \quad i = 1, 2, \dots, n$$

$$\rightarrow \mathbb{E}(\varepsilon_i) = 0, \text{Var}(\varepsilon_i) = \sigma^2, \text{ or } \varepsilon_i \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma^2)$$

参数估计 $\Rightarrow \hat{\beta}_1 = \frac{S_{XY}}{S_{XX}}, \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$

$$\hat{\sigma}^2 = \frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2 = \frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

总平方和分解公式: $\sum_{i=1}^n (y_i - \bar{y})^2 = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2$

决定系数: $r^2 = \frac{\sum_{i=1}^n (\hat{y}_i - \bar{y})^2}{\sum_{i=1}^n (y_i - \bar{y})^2} = \frac{\text{回归平方和}}{\text{总平方和}} = 1 - \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{\sum_{i=1}^n (y_i - \bar{y})^2}$

> summary(wages.lm)

Model : $\logwage = \beta_0 + \beta_1 \times education + \varepsilon$, $\varepsilon \sim N(0, \sigma^2)$
 or $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$, $\varepsilon_i \stackrel{i.i.d.}{\sim} (0, \sigma^2)$, $i = 1, 2, \dots, 2178$

Call:

lm(formula = logwage ~ education, data = wages)

Residuals: $e_i = y_i - \hat{y}_i = y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)$, $i = 1, 2, \dots, 2178$

Min	1Q	Median	3Q	Max
-1.78239	-0.25265	0.01636	0.27965	1.61101

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	1.239194	0.054974	22.54	<2e-16
education	0.078600	0.004262	18.44	<2e-16

$$2 \times P \left\{ t_{2176} > |22.54| \right\}$$

$$\frac{\hat{\beta}_0}{SE(\hat{\beta}_0)}$$

*** 显著性标志

$\hat{\beta}_1$

$$SE(\hat{\beta}_1)$$

$$\frac{\hat{\beta}_1}{SE(\hat{\beta}_1)}$$

$$2 \times P \left\{ t_{2176} > |18.44| \right\}$$

Signif. codes

' 0.001 '' 0.01

$\hat{\sigma} \leftarrow$ Residual standard error: 0.4038 on 2176 degrees of freedom $\rightarrow n - 2$

Multiple R-squared: 0.1351. Adjusted R-squared: 0.1347

F-statistic: 340 on 1 and 2176 DF, p-value: < 2.2e-16

$$R^2 = \frac{SSR}{SST} = 1 - \frac{SSE}{SST} = \widehat{Cor}(X, Y)^2$$

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

- 单因子方差分析 (ANOVA)

$$y_{ij} = \mu_j + \varepsilon_{ij} = (\mu + \alpha_j) + \varepsilon_{ij}, \quad i = 1, 2, \dots, m; \quad j = 1, 2, \dots, p$$

第 j 个水平的均值 \leftarrow μ_j \leftarrow $(\mu + \alpha_j)$ \leftarrow $\varepsilon_{ij} \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma^2)$

问题 $\begin{cases} H_0 : \mu_1 = \mu_2 = \dots = \mu_p = \mu \\ H_1 : \mu_k \neq \mu_l \text{ for some } k \text{ and } l \end{cases}$

- ▶ 单因子方差分析模型中的大部分信息都包含在方差分析表中
- ▶ 我们可以用 \mathcal{R} 中的 `aov()` 函数来作方差分析

Source	SS	df	MS	E(MS)
Treatment	$SS_{\text{treatment}} = m \sum_{j=1}^p (\bar{y}_j - \bar{y})^2$	$p - 1$	$\frac{SS_{\text{treatment}}}{p - 1}$	$\sigma^2 + \frac{m}{p - 1} \sum_{j=1}^p \alpha_j^2$
Error	$SS_{\text{full}} = \sum_{i=1}^m \sum_{j=1}^p (y_{ij} - \bar{y}_j)^2$	$p(m - 1)$	$\frac{SS_{\text{full}}}{p(m - 1)}$	σ^2

```

sales.aov <- aov(Sales ~ Strategy, data = Pullover)
summary(sales.aov)
  
```

```

> summary(sales.aov)
              Df Sum Sq Mean Sq F value Pr(>F)
Strategy      2  102.6   51.30   8.783 0.00115 **
Residuals    27  157.7    5.84
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
  
```

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

- 多重线性回归模型:

无截距模型:

$$\mathbf{y} = \mathcal{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$$

响应变量 Y , 解释变量 $\mathbf{X} = (X_1, X_2, \dots, X_p)^T$.

回归系数:

$$\boldsymbol{\beta} = \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{pmatrix}$$

误差项:

$$\boldsymbol{\varepsilon} = \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

参数估计 $\implies \hat{\boldsymbol{\beta}} = (\mathcal{X}^T \mathcal{X})^{-1} \mathcal{X}^T \mathbf{y}$

拟合值 $\implies \hat{\mathbf{y}} = \mathcal{X} \hat{\boldsymbol{\beta}} = \mathcal{X} (\mathcal{X}^T \mathcal{X})^{-1} \mathcal{X}^T \mathbf{y} = \mathcal{P} \mathbf{y}$ 向量 \mathbf{y} 在列空间 $C(\mathcal{X})$ 上的投影

残差 $\implies \mathbf{e} = \mathbf{y} - \hat{\mathbf{y}} = \mathbf{y} - \mathcal{P} \mathbf{y} = (\mathcal{I}_n - \mathcal{P}) \mathbf{y} = \mathcal{Q} \mathbf{y}$ 向量 \mathbf{y} 在 $C(\mathcal{X})$ 的正交补上的投影

决定系数:
$$r^2 = \frac{\sum_{i=1}^n (\hat{y}_i - \bar{y})^2}{\sum_{i=1}^n (y_i - \bar{y})^2} = \frac{\text{回归平方和}}{\text{总平方和}} = 1 - \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{\sum_{i=1}^n (y_i - \bar{y})^2}$$

调整后的决定系数:
$$r_{\text{adj}}^2 = r^2 - \frac{p(1 - r^2)}{n - p - 1}$$

已学知识点 (Recap)

第 3 章 多元中的一些概念与模型

- 多重线性回归模型:

```

> summary(pullover.lm2)

Call:
lm(formula = X1 ~ ., data = pullover)

Residuals:
    Min       1Q   Median       3Q      Max
-13.369  -9.406   1.599   5.151  19.729

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 65.66956   57.12507   1.150  0.29407
X2          -0.21578    0.32194  -0.670  0.52764
X3           0.48519    0.08678   5.591  0.00139 **
X4           0.84373    0.37400   2.256  0.06491 .
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 12.7 on 6 degrees of freedom
Multiple R-squared:  0.9067,    Adjusted R-squared:  0.8601
F-statistic: 19.44 on 3 and 6 DF,  p-value: 0.001713
  
```

$H_0 : \beta_1 = 0 \leftrightarrow H_1 : \beta_1 \neq 0 \implies$ 接受 H_0

$H_0 : \beta_2 = 0 \leftrightarrow H_1 : \beta_2 \neq 0 \implies$ 拒绝 H_0

$H_0 : \beta_3 = 0 \leftrightarrow H_1 : \beta_3 \neq 0 \implies$ 拒绝 H_0

$\left\{ \begin{array}{l} H_0 : \beta_1 = \beta_2 = \beta_3 = 0 \\ H_1 : \text{at least one } \beta_j \neq 0 \end{array} \right. \implies$ 拒绝 H_0

Multivariate Distributions

多元分布

概述

多元分布

分布函数与密度函数 (Distribution and Density Function)

矩与特征函数 (Moments and Characteristic Functions)

变量变换 (Transformations)

多元正态分布 (The Multinormal Distribution)

抽样分布与极限定理 (Sampling Distributions and Limit Theorems)

厚尾分布 (Heavy-Tailed Distributions)

自助法 (Bootstrap)

分布函数与密度函数

$$\text{随机向量 } \mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}$$

\mathbf{X} 的累积分布函数 (cumulative distribution function) 定义为

$$F(\mathbf{x}) = P(\mathbf{X} \leq \mathbf{x}) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_p \leq x_p)$$

▶ 对连续型随机变量 \mathbf{X} , 存在一个非负的概率密度函数 (probability density function) f 使得

$$F(\mathbf{x}) = \int_{-\infty}^{\mathbf{x}} f(\mathbf{u}) d\mathbf{u} = \int_{-\infty}^{x_p} \cdots \int_{-\infty}^{x_1} f(u_1, u_2, \dots, u_p) du_1 \cdots du_p$$

$$\int_{-\mathbb{R}^p} f(\mathbf{u}) d\mathbf{u} = 1, \quad f(\mathbf{x}) = \frac{\partial^p F(\mathbf{x})}{\partial x_1 \cdots \partial x_p}$$

分布函数与密度函数

$$\text{随机向量 } \mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}$$

\mathbf{X} 的累积分布函数 (cumulative distribution function) 定义为

$$F(\mathbf{x}) = P(\mathbf{X} \leq \mathbf{x}) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_p \leq x_p)$$

- ▶ 对离散型随机变量 \mathbf{X} , 我们关注该变量取值在可列个或有限个点 $\{c_j\}_{j \in J}$ 的概率,

$$P(\mathbf{X} \in D) = \sum_{\{j: c_j \in D\}} P(\mathbf{X} = c_j)$$

分布函数与密度函数

- 若将随机向量 $X = \begin{pmatrix} X_1 \\ \vdots \\ X_k \\ X_{k+1} \\ \vdots \\ X_p \end{pmatrix}$ 分块为 $X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}$, 其中 $X_1 \in \mathbb{R}^k$, $X_2 \in \mathbb{R}^{p-k}$, 则称

$$F_{X_1}(\mathbf{x}_1) = P(X_1 \leq \mathbf{x}_1) = F(x_1, \dots, x_k, \infty, \dots, \infty)$$

为**边缘 (marginal) 分布函数 (cdf)**. 称 $F = F(\mathbf{x})$ 为**联合 (joint) 分布函数 (cdf)**.

- ▶ 对连续型随机向量 X , $f_{X_1}(\mathbf{x}_1) = \int_{-\infty}^{\infty} f(\mathbf{x}_1, \mathbf{x}_2) d\mathbf{x}_2$.
- ▶ 给定 $X_1 = \mathbf{x}_1$ 时, X_2 的条件概率密度为 $f(\mathbf{x}_2 | \mathbf{x}_1) = \frac{f(\mathbf{x}_1, \mathbf{x}_2)}{f_{X_1}(\mathbf{x}_1)}$.

分布函数与密度函数

- 例：考虑函数

$$f(x_1, x_2) = \begin{cases} \frac{1}{2}x_1 + \frac{3}{2}x_2, & 0 \leq x_1, x_2 \leq 1 \\ 0, & \text{其它} \end{cases}$$

概率密度函数

$$f(x_1, x_2) \geq 0, \quad \iint f(x_1, x_2) dx_1 dx_2 = \frac{1}{2} \cdot \frac{x_1^2}{2} \Big|_0^1 + \frac{3}{2} \cdot \frac{x_2^2}{2} \Big|_0^1 = \frac{1}{4} + \frac{3}{4} = 1$$

边际密度：

$$\begin{cases} f_{X_1}(x_1) = \int f(x_1, x_2) dx_2 = \int_0^1 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_2 = \frac{1}{2}x_1 + \frac{3}{4}, & 0 \leq x_1 \leq 1 \\ f_{X_2}(x_2) = \int f(x_1, x_2) dx_1 = \int_0^1 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 = \frac{3}{2}x_2 + \frac{1}{4}, & 0 \leq x_2 \leq 1 \end{cases}$$

条件密度：

$$\begin{cases} f(x_2 | x_1) = \frac{f(x_1, x_2)}{f(x_1)} = \frac{\frac{1}{2}x_1 + \frac{3}{2}x_2}{\frac{1}{2}x_1 + \frac{3}{4}}, & 0 \leq x_2 \leq 1, (0 \leq x_1 \leq 1) \\ f(x_1 | x_2) = \frac{f(x_1, x_2)}{f(x_2)} = \frac{\frac{1}{2}x_1 + \frac{3}{2}x_2}{\frac{3}{2}x_2 + \frac{1}{4}}, & 0 \leq x_1 \leq 1, (0 \leq x_2 \leq 1) \end{cases}$$

分布函数与密度函数

定义 4.1 X_1 与 X_2 相互独立的充分必要条件是 $f(\mathbf{x}) = f(x_1, x_2) = f_{X_1}(x_1)f_{X_2}(x_2)$.

- ▶ X_1 与 X_2 相互独立当且仅当 $f(x_1 | x_2) = f_{X_1}(x_1)$, $f(x_2 | x_1) = f_{X_2}(x_2)$.
- ▶ 独立性可解释如下: 已知 $X_2 = x_2$ 不会改变 X_1 的概率, 反之亦然.



不同的联合概率密度函数可能会有相同的边际概率密度函数.

分布函数与密度函数

- 例：考虑下述两个概率密度函数

$$f(x_1, x_2) = \begin{cases} 1, & 0 < x_1, x_2 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_1, x_2) = \begin{cases} 1 + \alpha(2x_1 - 1)(2x_2 - 1), & 0 < x_1, x_2 < 1 \\ 0, & \text{其它} \end{cases}, \quad (-1 \leq \alpha \leq 1)$$

$$f_{X_1}(x_1) = \begin{cases} 1, & 0 < x_1 < 1 \\ 0, & \text{其它} \end{cases}, \quad f_{X_2}(x_2) = \begin{cases} 1, & 0 < x_2 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f_{X_1}(x_1) = \int_0^1 [1 + \alpha(2x_1 - 1)(2x_2 - 1)] dx_2 = 1, \quad 0 < x_1 < 1$$

$$f_{X_2}(x_2) = \int_0^1 [1 + \alpha(2x_1 - 1)(2x_2 - 1)] dx_1 = 1, \quad 0 < x_2 < 1$$

分布函数与密度函数

- 例：利用钞票数据集讨论独立性.

```

library(mclust)
data(banknote)
# 钞票数据的相关矩阵
round(cor(banknote[, 2:7]), digits = 2)
  
```

```

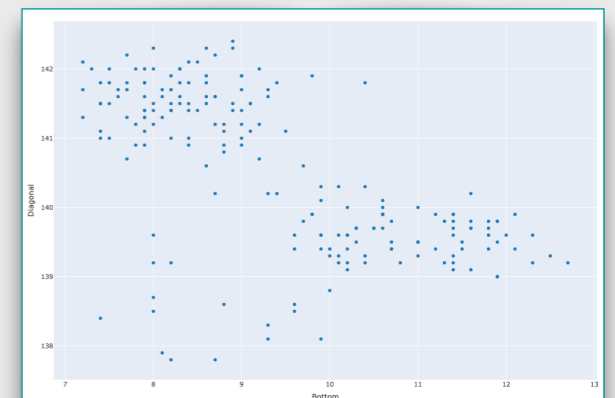
> # 钞票数据的相关矩阵
> round(cor(banknote[, 2:7]), digits = 2)
      Length Left Right Bottom Top Diagonal
Length  1.00  0.23  0.15 -0.19 -0.06  0.19
Left    0.23  1.00  0.74  0.41  0.36 -0.50
Right   0.15  0.74  1.00  0.49  0.40 -0.52
Bottom -0.19  0.41  0.49  1.00  0.14 -0.62
Top     -0.06  0.36  0.40  0.14  1.00 -0.59
Diagonal 0.19 -0.50 -0.52 -0.62 -0.59  1.00
  
```

X_4 与 X_6 几乎不独立.

X_4 与 X_6 的散点图

```

library(ggplot2)
library(plotly)
fig = plot_ly(banknote, x = ~Bottom, y = ~Diagonal, type = "scatter", mode = "markers") %>%
  layout(plot_bgcolor = '#e5ecf6',
         xaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'),
         yaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'))
fig
  
```



分布函数与密度函数

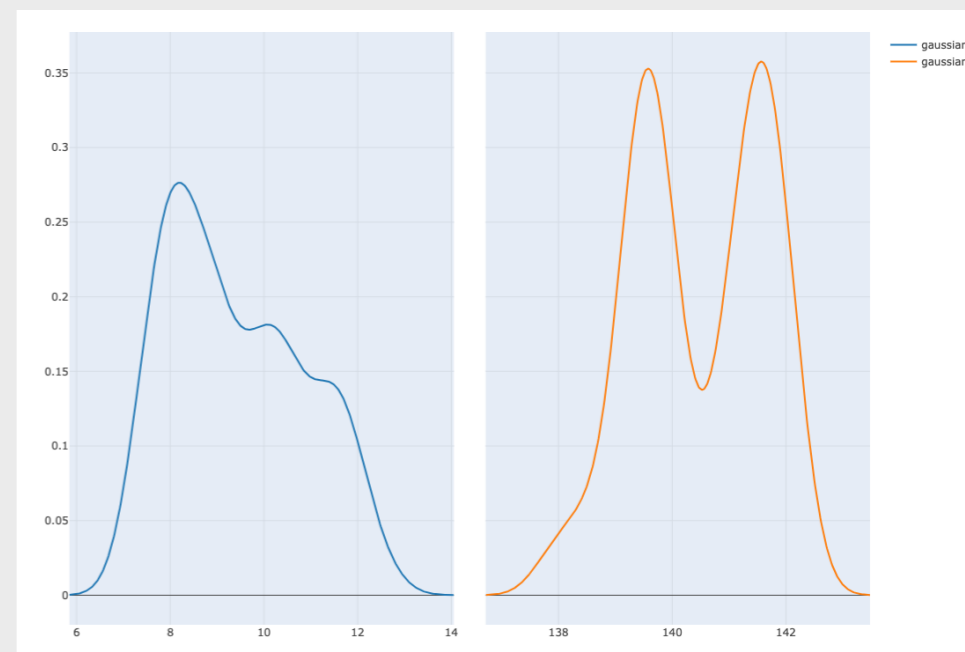
- 例：利用钞票数据集讨论独立性.

X4与X6的核密度图

```
fig1 = plot_ly()
dX4 = density(banknote$Bottom, kernel = "gaussian", na.rm = TRUE)
fig1 = add_lines(fig1, x = dX4$x, y = dX4$y, name = "gaussian") %>%
  layout(plot_bgcolor='#e5ecf6',
         xaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'),
         yaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'))
```

```
fig2 = plot_ly()
dX6 = density(banknote$Diagonal, kernel = "gaussian")
fig2 = add_lines(fig2, x = dX6$x, y = dX6$y, name = "gaussian") %>%
  layout(plot_bgcolor='#e5ecf6',
         xaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'),
         yaxis = list(zerolinecolor = '#fff', zerolinewidth = 2, gridcolor = 'fff'))
```

```
subplot(fig1, fig2, shareY = TRUE)
```



分布函数与密度函数

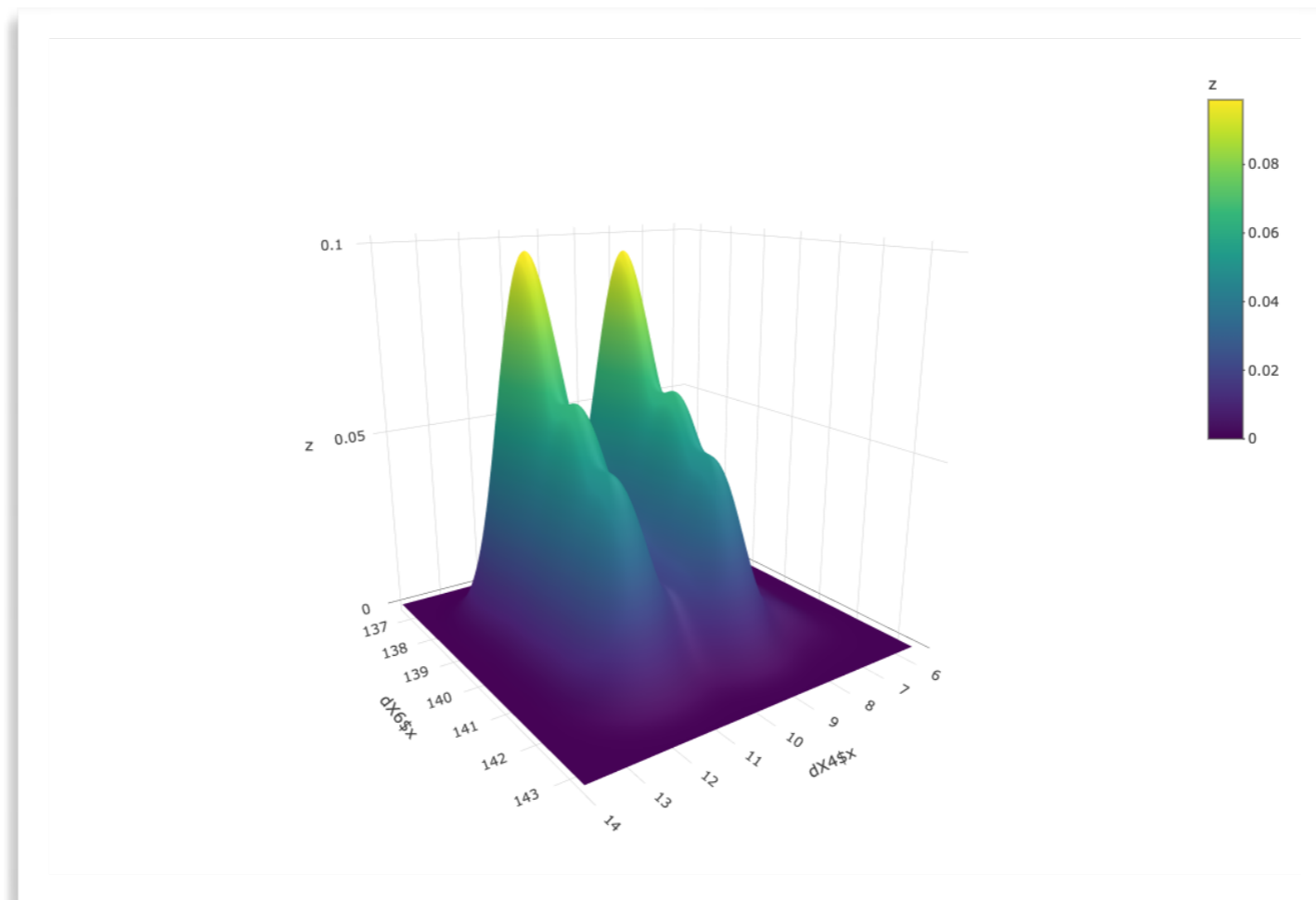
- 例：利用钞票数据集讨论独立性。
 - ▶ 如果 X_4 与 X_6 相互独立，则 $\hat{f}(x_4, x_6) = \hat{f}_{X_4} \cdot \hat{f}_{X_6}$ 。我们先作 $\hat{f}_{X_4} \cdot \hat{f}_{X_6}$ 的图形。

$\hat{f}_{X_4} \cdot \hat{f}_{X_6}$ 的图形

```
z = outer(dX4$y, dX6$y, function(a, b) a * b)
```

```
fig3 = plot_ly(x = ~dX4$x, y = ~dX6$x, z = ~z, type = "surface")
```

```
fig3
```

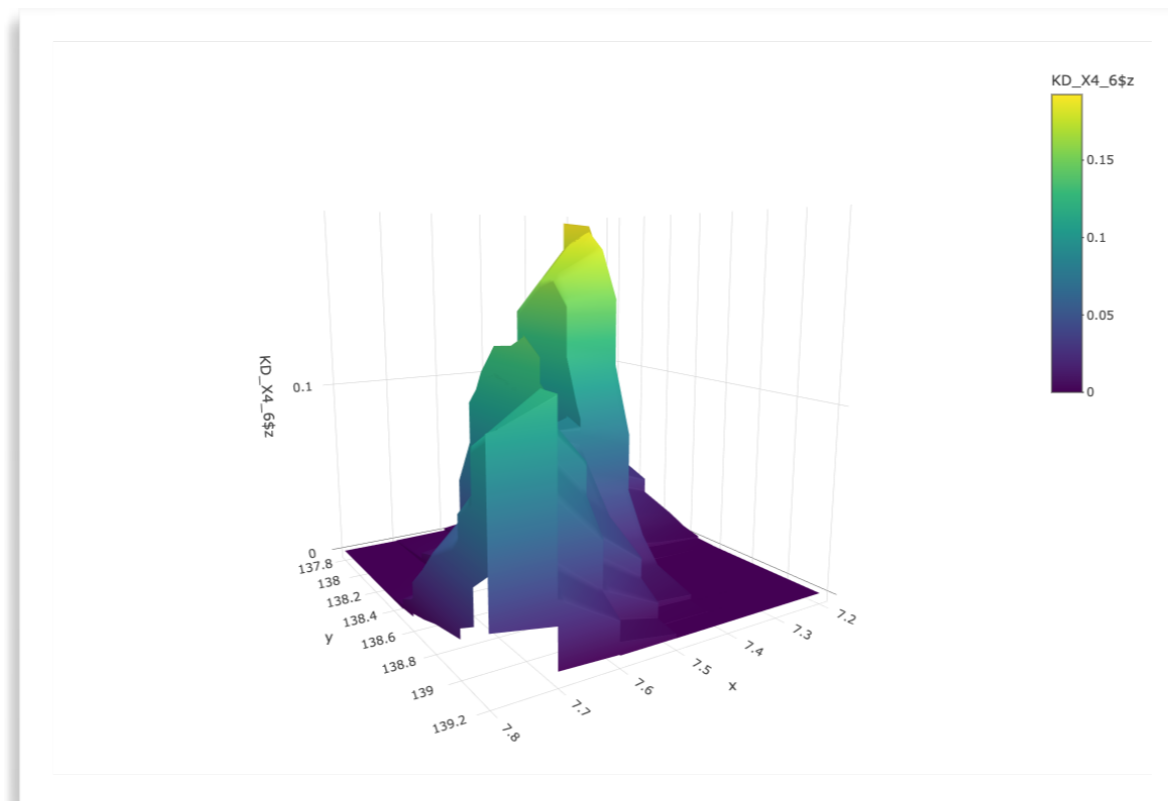


分布函数与密度函数

- 例：利用钞票数据集讨论独立性。
 - ▶ 再作 (X_4, X_6) 的联合密度函数的核密度图。

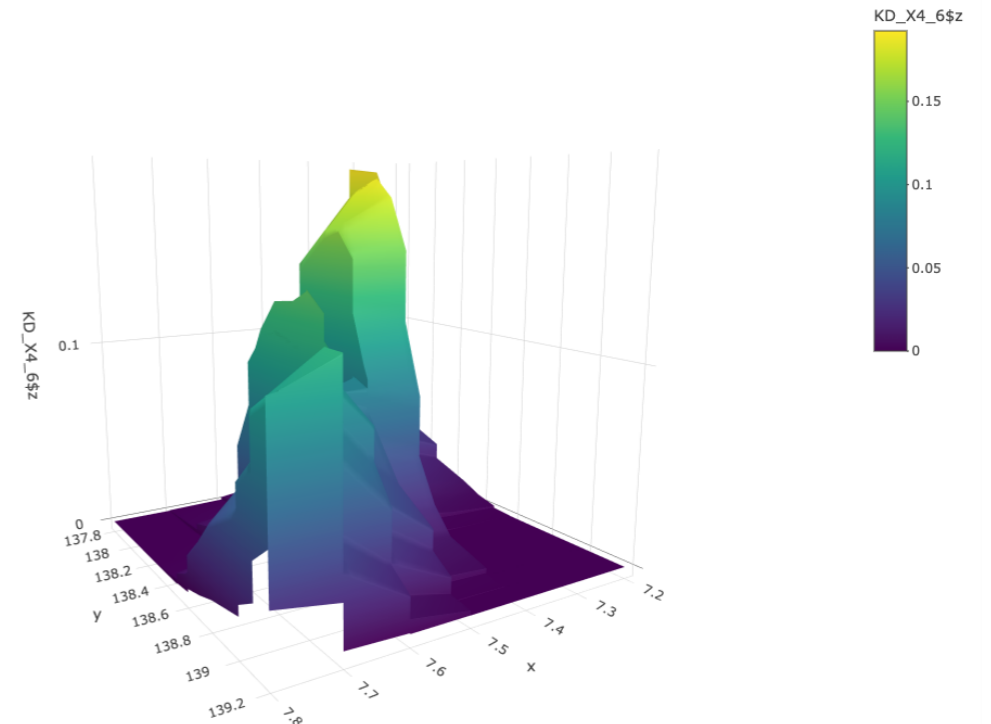
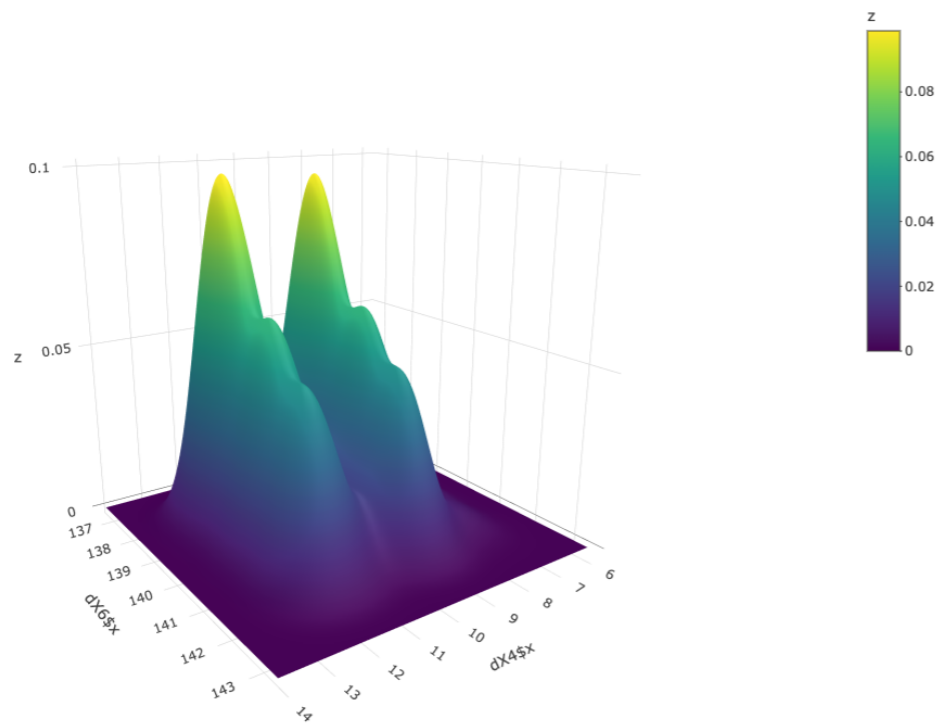
X4与X6的联合概率密度函数的核密度图

```
library(MASS)
x = sort(banknote$Bottom)
y = sort(banknote$Diagonal)
KD_X4_6 = kde2d(x, y, n = 25)
fig4 = plot_ly(x = ~x, y = ~y, z = ~KD_X4_6$z, type = "surface")
fig4
```



分布函数与密度函数

- 例：利用钞票数据集讨论独立性。
 - ▶ 如果 X_4 与 X_6 相互独立，则 $\hat{f}_{X_4} \cdot \hat{f}_{X_6}$ 与 (X_4, X_6) 的图形应该差异不大。



⇒ X_4 与 X_6 不独立

矩与特征函数 (Moments and Characteristic Functions)

- 矩 (moment) - 期望 (expectation) 与协方差矩阵 (covariance matrix)

▶ 期望：设随机向量 $\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}$ 的概率密度函数为 $f(\mathbf{x})$,

$$\mathbf{E}(\mathbf{X}) = \begin{pmatrix} \mathbf{E}(X_1) \\ \mathbf{E}(X_2) \\ \vdots \\ \mathbf{E}(X_p) \end{pmatrix} = \int \cdots \int \mathbf{x} f(\mathbf{x}) \, d\mathbf{x} = \begin{pmatrix} \int \cdots \int x_1 f(\mathbf{x}) \, d\mathbf{x} \\ \int \cdots \int x_2 f(\mathbf{x}) \, d\mathbf{x} \\ \vdots \\ \int \cdots \int x_p f(\mathbf{x}) \, d\mathbf{x} \end{pmatrix} \triangleq \boldsymbol{\mu}$$

$$\implies \mathbf{E}(\alpha \mathbf{X} + \beta \mathbf{Y}) = \alpha \mathbf{E}(\mathbf{X}) + \beta \mathbf{E}(\mathbf{Y})$$

$$\mathbf{E}(\mathcal{A}\mathbf{X}) = \mathcal{A}\mathbf{E}(\mathbf{X}) \quad \longleftarrow \quad \mathcal{A}_{q \times p} \text{ 为实矩阵}$$

$$X \text{ 与 } Y \text{ 相互独立} \implies \mathbf{E}(\mathbf{X}\mathbf{Y}^T) = \mathbf{E}(\mathbf{X})\mathbf{E}(\mathbf{Y}^T)$$

矩与特征函数 (Moments and Characteristic Functions)

- 矩 (moment) - 期望 (expectation) 与协方差矩阵 (covariance matrix) $\mu = E(X)$

▶ 协方差矩阵: $\text{Var}(X) = \Sigma = E \left[(X - \mu) (X - \mu)^T \right]$

▶ 若随机向量 X 的均值向量为 μ , 协方差矩阵为 Σ , 则可以将其表示为 $X \sim (\mu, \Sigma)$.

▶ 两个随机向量 $X \sim (\mu, \Sigma_{XX})$ 与 $Y \sim (\nu, \Sigma_{YY})$ 的协方差矩阵:

$$\Sigma_{XY} = \text{Cov}(X, Y) = E \left[(X - \mu) (Y - \nu)^T \right]$$

○ $\Sigma_{XY} = \Sigma_{YX}^T$

○ X 与 Y 相互独立 $\implies \text{Cov}(X, Y) = 0$.

○ $Z = \begin{pmatrix} X \\ Y \end{pmatrix} \implies \Sigma_{ZZ} = \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix}$

矩与特征函数 (Moments and Characteristic Functions)

- 矩 (moment) - 期望 (expectation) 与协方差矩阵 (covariance matrix)

- ▶ X 的一阶矩: $\mu = E(X)$.

- ▶ X 的二阶矩:

$$E\left(\mathbf{X}\mathbf{X}^T\right) = \left\{ E\left(X_i X_j\right) \right\}_{p \times p}, \quad i, j = 1, 2, \dots, p$$
$$= \begin{pmatrix} E\left(X_1^2\right) & E\left(X_1 X_2\right) & \cdots & E\left(X_1 X_p\right) \\ E\left(X_2 X_1\right) & E\left(X_2^2\right) & \cdots & E\left(X_2 X_p\right) \\ \vdots & \vdots & \ddots & \vdots \\ E\left(X_p X_1\right) & E\left(X_p X_2\right) & \cdots & E\left(X_p^2\right) \end{pmatrix}$$

矩与特征函数 (Moments and Characteristic Functions)

- 协方差矩阵 $\Sigma = \text{Var}(\mathbf{X})$ 的性质

$$\Sigma = \left(\sigma_{X_i X_j} \right)_{p \times p}, \quad \sigma_{X_i X_j} = \text{Cov}(X_i, X_j), \quad \sigma_{X_i X_i} = \text{Var}(X_i)$$

$$\Sigma = \text{E}(\mathbf{X}\mathbf{X}^T) - \boldsymbol{\mu}\boldsymbol{\mu}^T$$

$$\Sigma \geq 0$$

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix}$$

$$\boldsymbol{\mu} = \begin{pmatrix} \text{E}(X_1) \\ \text{E}(X_2) \\ \vdots \\ \text{E}(X_p) \end{pmatrix} = \boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{pmatrix}$$

- 方差与协方差的性质

$$\mathbf{a} = \begin{pmatrix} a_1 \\ a_2 \\ \vdots \\ a_p \end{pmatrix}$$

$$\text{Var}(\mathbf{a}^T \mathbf{X}) = \mathbf{a}^T \text{Var}(\mathbf{X}) \mathbf{a} = \sum_{i,j} a_i a_j \sigma_{X_i X_j}$$

$$\text{Var}(\mathcal{A} \mathbf{X} + \mathbf{b}) = \mathcal{A} \text{Var}(\mathbf{X}) \mathcal{A}^T$$

$$\text{Cov}(\mathbf{X} + \mathbf{Y}, \mathbf{Z}) = \text{Cov}(\mathbf{X}, \mathbf{Z}) + \text{Cov}(\mathbf{Y}, \mathbf{Z})$$

$$\text{Var}(\mathbf{X} + \mathbf{Y}) = \text{Var}(\mathbf{X}) + \text{Cov}(\mathbf{X}, \mathbf{Y}) + \text{Cov}(\mathbf{Y}, \mathbf{X}) + \text{Var}(\mathbf{Y})$$

$$\text{Cov}(\mathcal{A} \mathbf{X}, \mathcal{B} \mathbf{Y}) = \mathcal{A} \text{Cov}(\mathbf{X}, \mathbf{Y}) \mathcal{B}^T$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：设有概率密度函数 $f(x_1, x_2) = \begin{cases} \frac{1}{2}x_1 + \frac{3}{2}x_2, & 0 \leq x_1, x_2 \leq 1 \\ 0, & \text{otherwise} \end{cases}$

- ▶ 则其均值向量 $\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} = \begin{pmatrix} \frac{13}{24} \\ \frac{5}{8} \end{pmatrix}$

$$\mu_1 = \iint x_1 f(x_1, x_2) dx_1 dx_2 = \int_0^1 \int_0^1 x_1 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 dx_2 = \frac{13}{24}$$

$$\mu_2 = \iint x_2 f(x_1, x_2) dx_1 dx_2 = \int_0^1 \int_0^1 x_2 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 dx_2 = \frac{5}{8}$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：设有概率密度函数 $f(x_1, x_2) = \begin{cases} \frac{1}{2}x_1 + \frac{3}{2}x_2, & 0 \leq x_1, x_2 \leq 1 \\ 0, & \text{其它} \end{cases}$

▶ 则其均值向量 $\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} = \begin{pmatrix} \frac{13}{24} \\ \frac{5}{8} \end{pmatrix}$

▶ 其协方差矩阵 $\Sigma = \begin{pmatrix} \sigma_{X_1X_1} & \sigma_{X_1X_2} \\ \sigma_{X_2X_1} & \sigma_{X_2X_2} \end{pmatrix} = \begin{pmatrix} 0.0816 & -0.0052 \\ -0.0052 & 0.0677 \end{pmatrix}$

$$\sigma_{X_1X_1} = E(X_1^2) - \mu_1^2 = \int_0^1 \int_0^1 x_1^2 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 dx_2 - \left(\frac{13}{24} \right)^2 = \frac{3}{8} - \left(\frac{13}{24} \right)^2 \approx 0.0816$$

$$\sigma_{X_2X_2} = E(X_2^2) - \mu_2^2 = \int_0^1 \int_0^1 x_2^2 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 dx_2 - \left(\frac{5}{8} \right)^2 = \frac{11}{24} - \left(\frac{5}{8} \right)^2 \approx 0.0677$$

$$\sigma_{X_1X_2} = E(X_1X_2) - \mu_1\mu_2 = \int_0^1 \int_0^1 x_1x_2 \left(\frac{1}{2}x_1 + \frac{3}{2}x_2 \right) dx_1 dx_2 - \frac{13}{24} \times \frac{5}{8} = \frac{1}{3} - \frac{13}{24} \times \frac{5}{8} \approx -0.0052$$

矩与特征函数 (Moments and Characteristic Functions)

- 条件期望 (conditional expectation)
 - 给定 $X_1 = x_1$ 时, X_2 的条件位置参数
 - 给定 $X_2 = x_2$ 时, X_1 的条件位置参数

$$E(X_2 | x_1) \triangleq \int x_2 f(x_2 | x_1) dx_2, \quad E(X_1 | x_2) \triangleq \int x_1 f(x_1 | x_2) dx_1$$

$$\text{Var}(X_2 | X_1 = x_1) = E(X_2 X_2^T | X_1 = x_1) - E(X_2 | X_1 = x_1) E(X_2^T | X_1 = x_1)$$

$$\text{Var}(X) = E(XX^T) - \mu\mu^T$$


$$\text{Var}(X_1 | X_2 = x_2) = E(X_1 X_1^T | X_2 = x_2) - E(X_1 | X_2 = x_2) E(X_1^T | X_2 = x_2)$$

- 条件相关系数 (conditional correlations)

$$\rho_{X_2, X_3 | X_1 = x_1} = \frac{\text{Cov}(X_2, X_3 | X_1 = x_1)}{\sqrt{\text{Var}(X_2 | X_1 = x_1) \text{Var}(X_3 | X_1 = x_1)}}$$

给定 $X_1 = x_1$ 时, X_2 与 X_3 的偏相关系数 (partial correlation).

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑概率密度函数 $f(x_1, x_2, x_3) = \begin{cases} \frac{2}{3}(x_1 + x_2 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

 关于 x_1, x_2, x_3 对称


$$\Rightarrow f(x_1, x_2) = \int f(x_1, x_2, x_3) dx_3 = \begin{cases} \frac{2}{3} \left(x_1 + x_2 + \frac{1}{2} \right), & 0 < x_1, x_2 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_1, x_3) = \begin{cases} \frac{2}{3} \left(x_1 + x_3 + \frac{1}{2} \right), & 0 < x_1, x_3 < 1 \\ 0, & \text{其它} \end{cases}, \quad f(x_2, x_3) = \begin{cases} \frac{2}{3} \left(x_2 + x_3 + \frac{1}{2} \right), & 0 < x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

$$\Rightarrow f(x_1) = \int f(x_1, x_2) dx_2 = \begin{cases} \frac{2}{3} (x_1 + 1), & 0 < x_1 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_2) = \begin{cases} \frac{2}{3} (x_2 + 1), & 0 < x_2 < 1 \\ 0, & \text{其它} \end{cases}, \quad f(x_3) = \begin{cases} \frac{2}{3} (x_3 + 1), & 0 < x_3 < 1 \\ 0, & \text{其它} \end{cases}$$


矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑概率密度函数 $f(x_1, x_2, x_3) = \begin{cases} \frac{2}{3}(x_1 + x_2 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

 关于 x_1, x_2, x_3 对称

$$\Rightarrow f(x_1, x_2 | x_3) = \frac{f(x_1, x_2, x_3)}{f(x_3)} = \begin{cases} \frac{x_1 + x_2 + x_3}{x_3 + 1}, & 0 < x_1, x_2 < 1, \\ 0, & \text{其它,} \end{cases} \quad (0 < x_3 < 1)$$

$$\Rightarrow f(x_1 | x_3) = \frac{f(x_1, x_3)}{f(x_3)} = \begin{cases} \frac{x_1 + x_3 + \frac{1}{2}}{x_3 + 1}, & 0 < x_1 < 1 \\ 0, & \text{其它} \end{cases}, \quad (0 < x_3 < 1)$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑概率密度函数 $f(x_1, x_2, x_3) = \begin{cases} \frac{2}{3}(x_1 + x_2 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

 关于 x_1, x_2, x_3 对称

► 计算可得以下各阶矩


$$E(X_i) = \frac{5}{9}, \quad E(X_i^2) = \frac{7}{18}; \quad E(X_i X_j) = \frac{11}{36}, \quad (i \neq j)$$

$$\text{Var}(X_i) = E(X_i^2) - [E(X_i)]^2 = \frac{7}{18} - \left(\frac{5}{9}\right)^2 = \frac{13}{162}$$

$$\text{Cov}(X_i, X_j) = E(X_i X_j) - E(X_i) E(X_j) = \frac{11}{36} - \left(\frac{5}{9}\right)^2 = -\frac{1}{324}$$

$$\Rightarrow \Sigma = \begin{pmatrix} \frac{13}{162} & -\frac{1}{324} & -\frac{1}{324} \\ \frac{1}{324} & \frac{13}{162} & -\frac{1}{324} \\ \frac{1}{324} & -\frac{1}{324} & \frac{13}{162} \end{pmatrix} \Rightarrow \rho_{X_1 X_2} = \frac{-\frac{1}{324}}{\sqrt{\frac{13}{162} \cdot \frac{13}{162}}} = -\frac{1}{26} \approx -0.0385$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑概率密度函数 $f(x_1, x_2, x_3) = \begin{cases} \frac{2}{3}(x_1 + x_2 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$


▶ 计算可得以下各阶矩

$$E(X_1 | X_3 = x_3) = E(X_2 | X_3 = x_3) = \frac{1}{12} \left(\frac{6x_3 + 7}{x_3 + 1} \right)$$

$$E(X_1^2 | X_3 = x_3) = E(X_2^2 | X_3 = x_3) = \frac{1}{12} \left(\frac{4x_3 + 5}{x_3 + 1} \right)$$

$$E(X_1 X_2 | X_3 = x_3) = \frac{1}{12} \left(\frac{3x_3 + 4}{x_3 + 1} \right)$$

不是 x_3 的线性函数



矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑概率密度函数 $f(x_1, x_2, x_3) = \begin{cases} \frac{2}{3}(x_1 + x_2 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$
→ 关于 x_1, x_2, x_3 对称

$$\text{Var}(X_1 | X_2 = x_2) = E(X_1 X_1^T | X_2 = x_2) - E(X_1 | X_2 = x_2) E(X_1^T | X_2 = x_2)$$

$$\Rightarrow \text{Var} \left(\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \middle| X_3 = x_3 \right) = \begin{pmatrix} \frac{12x_3^2 + 24x_3 + 11}{144(x_3 + 1)^2} & \frac{-1}{144(x_3 + 1)^2} \\ \frac{-1}{144(x_3 + 1)^2} & \frac{12x_3^2 + 24x_3 + 11}{144(x_3 + 1)^2} \end{pmatrix}$$

$$\text{Cov}(X_2, X_3 | X_1 = x_1)$$

$$\rho_{X_2, X_3 | X_1 = x_1} = \frac{\text{Cov}(X_2, X_3 | X_1 = x_1)}{\sqrt{\text{Var}(X_2 | X_1 = x_1) \text{Var}(X_3 | X_1 = x_1)}}$$

$$\frac{-1}{144(x_3 + 1)^2}$$

$$\Rightarrow \rho_{X_1 X_2 | X_3 = x_3} = \frac{\frac{-1}{144(x_3 + 1)^2}}{\sqrt{\frac{12x_3^2 + 24x_3 + 11}{144(x_3 + 1)^2} \cdot \frac{12x_3^2 + 24x_3 + 11}{144(x_3 + 1)^2}}}$$

$x_3 \in (0, 1)$

$$\rho_{X_1 X_2 | X_3 = x_3} \in \left(-\frac{1}{11}, -\frac{1}{47} \right) \approx (-0.0909, -0.0213)$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑密度函数 $f(x_1, x_2, x_3) = \begin{cases} 2x_2(x_1 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

X_2 与 (X_1, X_3) 独立

关于 x_1 与 x_3 对称

$$\Rightarrow f(x_1, x_3) = \int f(x_1, x_2, x_3) dx_2 = \begin{cases} x_1 + x_3, & 0 < x_1, x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_2) = \begin{cases} 2x_2, & 0 < x_2 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_1) = \int f(x_1, x_3) dx_3 = \begin{cases} x_1 + \frac{1}{2}, & 0 < x_1 < 1 \\ 0, & \text{其它} \end{cases}, \quad f(x_3) = \int f(x_1, x_3) dx_1 = \begin{cases} x_3 + \frac{1}{2}, & 0 < x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

$$\Rightarrow E(\mathbf{X}) = \begin{pmatrix} \frac{7}{12} \\ \frac{2}{3} \\ \frac{7}{12} \end{pmatrix}, \quad \Sigma = \begin{pmatrix} \frac{1}{144} & 0 & -\frac{1}{144} \\ 0 & \frac{1}{18} & 0 \\ -\frac{1}{144} & 0 & \frac{11}{144} \end{pmatrix}$$

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑密度函数 $f(x_1, x_2, x_3) = \begin{cases} 2x_2(x_1 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

X_2 与 (X_1, X_3) 独立

关于 x_1 与 x_3 对称

- 讨论给定 $X_3 = x_3$ 时 (X_1, X_2) 的条件分布.

$$f(x_3) = \begin{cases} x_3 + \frac{1}{2}, & 0 < x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_1, x_2 | x_3) = \frac{f(x_1, x_2, x_3)}{f(x_3)} = \begin{cases} \frac{4(x_1 + x_3)x_2}{2x_3 + 1}, & 0 < x_1, x_2 < 1, \quad (0 < x_3 < 1) \\ 0, & \text{其它} \end{cases}$$

$$f(x_1, x_3) = \begin{cases} x_1 + x_3, & 0 < x_1, x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

$$f(x_1 | x_3) = \frac{f(x_1, x_3)}{f(x_3)} = \begin{cases} \frac{2(x_1 + x_3)}{2x_3 + 1}, & 0 < x_1 < 1, \quad (0 < x_3 < 1) \\ 0, & \text{其它} \end{cases}$$

$$f(x_2 | x_3) = f(x_2) = 2x_2, \quad 0 < x_2 < 1 \quad (0 < x_3 < 1)$$

$$\Rightarrow f(x_1, x_2 | x_3) = f(x_1 | x_3) \cdot f(x_2 | x_3)$$

给定 $X_3 = x_3$ 时, X_1 与 X_2 的条件分布相互独立

矩与特征函数 (Moments and Characteristic Functions)

- 例：考虑密度函数 $f(x_1, x_2, x_3) = \begin{cases} 2x_2(x_1 + x_3), & 0 < x_1, x_2, x_3 < 1 \\ 0, & \text{其它} \end{cases}$

X_2 与 (X_1, X_3) 独立

关于 x_1 与 x_3 对称

$$f(x_3) = \begin{cases} x_3 + \frac{1}{2}, & 0 < x_3 < 1 \\ 0, & \text{其它} \end{cases}$$

- 讨论给定 $X_3 = x_3$ 时 (X_1, X_2) 的条件分布.

$$f(x_1, x_2 | x_3) = \frac{f(x_1, x_2, x_3)}{f(x_3)} = \begin{cases} \frac{4(x_1 + x_3)x_2}{2x_3 + 1}, & 0 < x_1, x_2 < 1, \quad (0 < x_3 < 1) \\ 0, & \text{其它} \end{cases}$$

$$\Rightarrow E \left[\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \middle| X_3 = x_3 \right] = \begin{pmatrix} \frac{1}{3} \left(\frac{2 + 3x_3}{1 + 2x_3} \right) \\ \frac{2}{3} \end{pmatrix}, \quad (0 < x_3 < 1)$$

$$\Rightarrow \text{Var} \left[\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \middle| X_3 = x_3 \right] = \begin{pmatrix} \frac{1}{18} \left(\frac{6x_3^2 + 6x_3 + 1}{(2x_3 + 1)^2} \right) & 0 \\ 0 & \frac{1}{18} \end{pmatrix}, \quad (0 < x_3 < 1)$$

矩与特征函数 (Moments and Characteristic Functions)

- 条件期望的性质

$$E\left(X_2 \mid X_1 = x_1\right) = h\left(x_1\right), \quad \text{Var}\left(X_2 \mid X_1 = x_1\right) = g\left(x_1\right)$$

- ▶ 利用条件期望定义随机变量

$$h\left(X_1\right) \triangleq E\left(X_2 \mid X_1\right), \quad g\left(X_1\right) \triangleq \text{Var}\left(X_2 \mid X_1\right)$$

- ▶ 有趣的性质

$$E\left(X_2\right) = E\left[E\left(X_2 \mid X_1\right)\right]$$

$$\text{Var}\left(X_2\right) = E\left[\text{Var}\left(X_2 \mid X_1\right)\right] + \text{Var}\left[E\left(X_2 \mid X_1\right)\right]$$

矩与特征函数 (Moments and Characteristic Functions)

- 例:** 考虑密度函数 $f(x_1, x_2) = \begin{cases} 2e^{-\frac{x_2}{x_1}}, & 0 < x_1 < 1, x_2 > 0 \\ 0, & \text{其它} \end{cases}$

$$\Rightarrow f(x_1) = \int f(x_1, x_2) dx_2 = \begin{cases} 2x_1, & 0 < x_1 < 1 \\ 0, & \text{其它} \end{cases} \Rightarrow E(X_1) = \frac{2}{3}, \quad \text{Var}(X_1) = \frac{1}{18}$$

$$f(x_2 | x_1) = \frac{f(x_1, x_2)}{f(x_1)} = \begin{cases} \frac{1}{x_1} e^{-\frac{x_2}{x_1}}, & x_2 > 0 \\ 0, & \text{其它} \end{cases}, \quad (0 < x_1 < 1)$$


$$\Rightarrow E(X_2 | X_1) = X_1, \quad \text{Var}(X_2 | X_1) = X_1^2$$

- 无须计算 $f(x_2)$ 的表达式, 我们就可以得到

$$E(X_2) = E\left[E(X_2 | X_1)\right] = E(X_1) = \frac{2}{3}$$

$$\text{Var}(X_2) = E\left[\text{Var}(X_2 | X_1)\right] + \text{Var}\left[E(X_2 | X_1)\right] = E(X_1^2) + \text{Var}(X_1) = \frac{2}{4} + \frac{1}{18} = \frac{10}{18}$$

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- 条件期望 $E(X_2 | X_1)$ 看作 X_1 的一个函数 $h(X_1)$, 可以解释为通过 X_1 的一个函数对 X_2 的一种条件近似.
  称之为 X_2 关于 X_1 的回归函数
- ▶ 该近似的误差项为 $U = X_2 - E(X_2 | X_1)$

定理 4.3 设 $X_1 \in \mathbb{R}^k$, $X_2 \in \mathbb{R}^{p-k}$, $U = X_2 - E(X_2 | X_1)$. 则

1. $E(U) = \mathbf{0}$.
2. 用 X_1 的一个函数 $h(X_1)$ 对 X_2 作近似, 则 $E(X_2 | X_1)$ 是一个“最佳”逼近, 其中 $h: \mathbb{R}^k \rightarrow \mathbb{R}^{p-k}$. 所谓“最佳”是指均方误差 (MSE) 最小的意义下, 其中

$$\text{MSE}(h) = E \left\{ \left[X_2 - h(X_1) \right]^T \left[X_2 - h(X_1) \right] \right\}.$$